## Discriminant Analysis

## Discriminant analysis

- ANOVA and MANOVA: predict a (counted/measured) response from group membership.

Discriminant analysis: predict group membership based on counted/measured variables.

- Covers same ground as logistic regression (and its variations), but emphasis on classifying observed data into correct groups.
- Does so by searching for linear combination of original variables that best separates data into groups (canonical variables).
$\rightarrow$ Assumption here that groups are known (for data we have). If trying to "best separate" data into unknown groups, see cluster analysis.


## Packages

```
library(MASS, exclude = "select")
library(tidyverse)
library(ggrepel)
library(ggbiplot)
library(MVTests) # for Box M test
library(conflicted)
conflict_prefer("arrange", "dplyr")
conflict_prefer("summarize", "dplyr")
conflict_prefer("select", "dplyr")
conflict_prefer("filter", "dplyr")
conflict_prefer("mutate", "dplyr")
```

- ggrepel allows labelling points on a plot so they don't overwrite each other.
ggbiplot uses plyr rather than dplyr, which has functions by similar names.


## About select

Both dplyr (in tidyverse) and MASS have a function called select, and they do different things.

- How do you know which select is going to get called?
- With library, the one loaded last is visible, and others are not.
- Thus we can access the select in dplyr but not the one in MASS. If we wanted that one, we'd have to say MASS: :select.
- Better: load conflicted package. Any time you load two packages containing functions with same name, you get error and have to choose between them.


## Example 1: seed yields and weights

```
my_url <- "http://ritsokiguess.site/datafiles/manova1.txt"
hilo <- read_delim(my_url, " ")
g <- ggplot(hilo, aes(x = yield, y = weight,
    colour = fertilizer)) + geom_point(size = 4)
```

Recall data from MANOVA: needed a multivariate analysis to find difference in seed yield and weight based on whether they were high or low fertilizer.


## Basic discriminant analysis

hilo. 1 <- lda(fertilizer ~ yield + weight, data = hilo)

- Uses lda from package MASS.
- "Predicting" group membership from measured variables.


## Output

```
hilo.1
Call:
lda(fertilizer ~ yield + weight, data = hilo)
Prior probabilities of groups:
high low
    0.5 0.5
Group means:
    yield weight
high 35.0 13.25
low 32.5 12.00
Coefficients of linear discriminants:
    LD1
yield -0.7666761
weight -1.2513563
```


## Things to take from output

- Group means: high-fertilizer plants have (slightly) higher mean yield and weight than low-fertilizer plants.
- "Coefficients of linear discriminants": LD1, LD2, ...are scores constructed from observed variables that best separate the groups.

For any plant, get LD1 score by taking -0.76 times yield plus -1.25 times weight, add up, standardize.
the LD1 coefficients are like slopes:

- if yield higher, LD1 score for a plant lower
- if weight higher, LD1 score for a plant lower
- High-fertilizer plants have higher yield and weight, thus low (negative) LD1 score. Low-fertilizer plants have low yield and weight, thus high (positive) LD1 score.
- One LD1 score for each observation. Plot with actual groups.


## How many linear discriminants?

- Smaller of these:
- Number of variables
- Number of groups minus 1
- Seed yield and weight: 2 variables, 2 groups, $\min (2,2-1)=1$.


## Getting LD scores

Feed output from LDA into predict:

```
p <- predict(hilo.1)
hilo.2 <- cbind(hilo, p)
hilo.2
```

|  | fertilizer yield weight |  |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| class | posterior.high | posterior.low | LD1 |  |  |  |  |
| 1 | low | 34 | 10 | low | $2.108619 \mathrm{e}-05$ | $9.999789 \mathrm{e}-01$ | 3.0931414 |
| 2 | low | 29 | 14 | low | $1.245320 \mathrm{e}-03$ | $9.987547 \mathrm{e}-01$ | 1.9210963 |
| 3 | low | 35 | 11 | low | $2.315016 \mathrm{e}-02$ | $9.768498 \mathrm{e}-01$ | 1.0751090 |
| 4 | low | 32 | 13 | low | $4.579036 \mathrm{e}-02$ | $9.542096 \mathrm{e}-01$ | 0.8724245 |
| 5 | high | 33 | 14 | high | $9.817958 \mathrm{e}-01$ | $1.820422 \mathrm{e}-02$ | -1.1456079 |
| 6 | high | 38 | 12 | high | $9.998195 \mathrm{e}-01$ | $1.804941 \mathrm{e}-04$ | -2.4762756 |
| 7 | high | 34 | 13 | high | $9.089278 \mathrm{e}-01$ | $9.107216 \mathrm{e}-02$ | -0.6609276 |
| 8 | high | 35 | 14 | high | $9.999109 \mathrm{e}-01$ | $8.914534 \mathrm{e}-05$ | -2.6789600 |

## LD1 scores in order

Most positive LD1 score is most obviously low fertilizer, most negative is most obviously high:

```
hilo.2 %%% select(fertilizer, yield, weight, LD1) %>%
    arrange(desc(LD1))
```

    fertilizer yield weight LD1
    1 low $34 \quad 10 \quad 3.0931414$
2 low 29141.9210963
3 low $35 \quad 111.0751090$
4 low $32 \quad 13 \quad 0.8724245$
7 high $34 \quad 13-0.6609276$
5 high $33 \quad 14-1.1456079$
6 high $38 \quad 12$-2.4762756
8 high $35 \quad 14-2.6789600$

High fertilizer have yield and weight high, negative LD1 scores.

## Plotting LD1 scores

With one LD score, plot against (true) groups, eg. boxplot: ggplot(hilo.2, aes(x = fertilizer, y = LD1)) + geom_boxplot


fertilizer

## What else is in hilo. 2?

- class: predicted fertilizer level (based on values of yield and weight).
posterior: predicted probability of being low or high fertilizer given yield and weight.
$\rightarrow$ LD1: scores for (each) linear discriminant (here is only LD1) on each observation.


## Predictions and predicted groups

```
...based on yield and weight:
hilo.2 %>% select(yield, weight, fertilizer, class)
    yield weight fertilizer class
13410 low low
22914 low low
\(3 \quad 35 \quad 11\) low low
43213 low low
\(5 \quad 33 \quad 14\) high high
\(6 \quad 38 \quad 12 \quad\) high high
\begin{tabular}{llll}
7 & 34 & 13 & high high
\end{tabular}
\(8 \quad 35 \quad 14\) high high
```


## Count up correct and incorrect classificationot()

```
with(hilo.2, table(obs = fertilizer, pred = class))
```

| pred |  |  |
| :---: | ---: | ---: |
| obs | high | low |
| high | 4 | 0 |
| low | 0 | 4 |

- Each predicted fertilizer level is exactly same as observed one (perfect prediction).
- Table shows no errors: all values on top-left to bottom-right diagonal.


## Posterior probabilities

show how clear-cut the classification decisions were:

```
hilo.2 %>%
    mutate(across(starts_with("posterior"), \(p) round(p, 4))) %>%
    select(-LD1)
```

    fertilizer yield weight class posterior.high posterior.low
    | 1 | low | 34 | 10 | low | 0.0000 | 1.0000 |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
| 2 | low | 29 | 14 | low | 0.0012 | 0.9988 |
| 3 | low | 35 | 11 | low | 0.0232 | 0.9768 |
| 4 | low | 32 | 13 | low | 0.0458 | 0.9542 |
| 5 | high | 33 | 14 | high | 0.9818 | 0.0182 |
| 6 | high | 38 | 12 | high | 0.9998 | 0.0002 |
| 7 | high | 34 | 13 | high | 0.9089 | 0.0911 |
| 8 | high | 35 | 14 | high | 0.9999 | 0.0001 |

Only obs. 7 has any doubt: yield low for a high-fertilizer, but high weight makes up for it.

## Example 2: the peanuts

```
my_url <- "http://ritsokiguess.site/datafiles/peanuts.txt"
peanuts <- read_delim(my_url, " ")
peanuts
# A tibble: 12 x 6
            obs location variety y smk w
    <dbl> <dbl> <dbl> <dbl> <dbl> <dbl>
    1 1 1 5 5 195. 153. 51.4
    2 2 1 5 5 194. 168. 53.7
    3 3 2 2 5 190. 140. 55.5
    4 4 2 5 5 180. 121. 44.4
    5 5 1 1 % 6 203 157. 49.8
    6 6 1 % 6 196. 166 45.8
    7 7 2 2 6 203. 166. 60.4
    8 8 2 6 198. 162. 54.1
    9 9 1 % 8 194. 164. 57.8
```



```
11 11 2 8 202. 167. 65
12 12 2 % 8 200 174. 67.2
```

$\rightarrow$ Recall: location and variety both significant in MANOVA. Make combo of them (over):

## Location-variety combos

```
peanuts %>%
    unite(combo, c(variety, location)) -> peanuts.combo
peanuts.combo
# A tibble: 12 x 5
    obs combo y smk W
    1 5_1 195. 153. 51.4
    2 2 5_1 194. 168. 53.7
    3 3 5_2 190. 140. 55.5
    4 4 5_2 180. 121. 44.4
    5 5 6_1 203 157. 49.8
    6 6 6_1 196. 166 45.8
    7 7 6_2 203. 166. 60.4
    8 8 6_2 198. 162. 54.1
    9 9 8_1 194. 164. 57.8
10 10 8_1 187 165. 58.6
11 11 8_2 202. 167. 65
12 12 8_2 200 174. 67.2
```


## Discriminant analysis

```
# peanuts.1 <- lda(str_c(location, variety, sep = "_") ~ y + smk + w, data = peanuts)
peanuts.1 <- lda(combo ~ y + smk + w, data = peanuts.combo)
peanuts.1
Call:
lda(combo ~ y + smk + w, data = peanuts.combo)
Prior probabilities of groups:
    5_1 5_2 6_1 6_2 8_1 
0.1666667 0.1666667 0.1666667 0.1666667 0.1666667 0.1666667
Group means:
    y smk w
5_1 194.80 160.40 52.55
5_2 185.05 130.30 49.95
6_1 199.45 161.40 47.80
6_2 200.15 163.95 57.25
8_1 190.25 164.80 58.20
8_2 200.75 170.30 66.10
Coefficients of linear discriminants:
    LD1 LD2 LD3
y 0.4027356 0.02967881 0.18839237
smk 0.1727459 -0.06794271 -0.09386294
w -0.5792456 -0.16300221 0.07341123
Proportion of trace:
    LD1 LD2 LD3
0.8424 0.1317 0.0258
```


## Comments

- Now 3 LDs (3 variables, 6 groups, $\min (3,6-1)=3$ ).
$\rightarrow$ Relationship of LDs to original variables. Look for coeffs far from zero:
peanuts. $1 \$$ scaling

|  | LD1 | LD2 | LD3 |
| :--- | ---: | ---: | ---: |
| y | 0.4027356 | 0.02967881 | 0.18839237 |
| smk | 0.1727459 | -0.06794271 | -0.09386294 |
| w | -0.5792456 | -0.16300221 | 0.07341123 |

- high LD1 mainly high y or low w.
- high LD2 mainly low w.
- Proportion of trace values show relative importance of LDs: LD1 much more important than LD2; LD3 worthless.


## The predictions and misclassification

p <- predict (peanuts.1)
peanuts. 2 <- cbind(peanuts.combo, p)
peanuts. 2

|  | obs | combo | y | smk | w |  |  |  |  |  |  | class | posterior.5_1 | posteriol |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | 1 | $5 \_1$ | 195.3 | 153.1 | 51.4 | $5 \_1$ | $6.862288 \mathrm{e}-01$ | 1.825787 |  |  |  |  |  |  |
| 2 | 2 | $5 \_1$ | 194.3 | 167.7 | 53.7 | $5 \_1$ | $7.269338 \mathrm{e}-01$ | 7.55585 |  |  |  |  |  |  |
| 3 | 3 | $5 \_2$ | 189.7 | 139.5 | 55.5 | $5 \_2$ | $1.624097 \mathrm{e}-12$ | 9.99635 |  |  |  |  |  |  |
| 4 | 4 | $5 \_2$ | 180.4 | 121.1 | 44.4 | $5 \_2$ | $1.702156 \mathrm{e}-16$ | 1.00000 |  |  |  |  |  |  |
| 5 | 5 | $6 \_1$ | 203.0 | 156.8 | 49.8 | $6 \_1$ | $4.262552 \mathrm{e}-05$ | 1.50008 |  |  |  |  |  |  |
| 6 | 6 | $6 \_1$ | 195.9 | 166.0 | 45.8 | $6 \_1$ | $9.681355 \mathrm{e}-07$ | 1.07119 |  |  |  |  |  |  |
| 7 | 7 | $6 \_2$ | 202.7 | 166.1 | 60.4 | $6 \_2$ | $1.324922 \mathrm{e}-01$ | $5.98906!$ |  |  |  |  |  |  |
| 8 | 8 | $6 \_2$ | 197.6 | 161.8 | 54.1 | $5 \_1$ | $5.286987 \mathrm{e}-01$ | 2.03799 |  |  |  |  |  |  |
| 9 | 9 | $8 \_1$ | 193.5 | 164.5 | 57.8 | $8 \_1$ | $2.298649 \mathrm{e}-02$ | 6.924748 |  |  |  |  |  |  |
| 10 | 10 | $8 \_1$ | 187.0 | 165.1 | 58.6 | $8 \_1$ | $1.572134 \mathrm{e}-08$ | 5.77368 |  |  |  |  |  |  |
| 11 | 11 | $8 \_2$ | 201.5 | 166.8 | 65.0 | $8 \_2$ | $8.160707 \mathrm{e}-05$ | 6.48149 |  |  |  |  |  |  |
| 12 | 12 | $8 \_2$ | 200.0 | 173.8 | 67.2 | $8 \_2$ | $1.509768 \mathrm{e}-06$ | 1.55714 |  |  |  |  |  |  | posterior.6_2 posterior.8_1 posterior.8_2 x.LD1

## Posterinr nrohahilities

```
peanuts.2 %>%
    mutate(across(starts_with("posterior"), \(p) round(p, 2))) %>%
    select(combo, class, starts_with("posterior"))
```

    combo class posterior.5_1 posterior.5_2 posterior.6_1 posterior.6_2
    | 1 | 5_1 | 5_1 | 0.69 | 0 | 0 | 0.31 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 2 | 5_1 | 5_1 | 0.73 | 0 | 0 | 0.27 |
| 3 | 5_2 | 5_2 | 0.00 | 1 | 0 | 0.00 |
| 4 | 5_2 | 5_2 | 0.00 | 1 | 0 | 0.00 |
| 5 | 6_1 | 6_1 | 0.00 | 0 | 1 | 0.00 |
| 6 | 6_1 | 6_1 | 0.00 | 0 | 1 | 0.00 |
| 7 | 6_2 | 6_2 | 0.13 | 0 | 0 | 0.87 |
| 8 | 6_2 | 5_1 | 0.53 | 0 | 0 | 0.47 |
| 9 | 8_1 | 8_1 | 0.02 | 0 | 0 | 0.02 |
| 10 | 8_1 | 8_1 | 0.00 | 0 | 0 | 0.00 |
| 11 | 8_2 | 8_2 | 0.00 | 0 | 0 | 0.00 |
| 12 | 8_2 | 8_2 | 0.00 | 0 | 0 | 0.00 |

    posterior.8_1 posterior.8_2
            \(0.00 \quad 0.00\)
            \(0.00 \quad 0.00\)
            \(0.00 \quad 0.00\)
            \(0.00 \quad 0.00\)
            \(0.00 \quad 0.00\)
            \(0.00 \quad 0.00\)
            \(0.00 \quad 0.00\)
            \(\begin{array}{ll}0.00 & 0.00\end{array}\)
    
## Discriminant scores, again

- How are discriminant scores related to original variables?
- Construct data frame with original data and discriminant scores side by side:

```
peanuts.1$scaling
```

|  | LD1 | LD2 | LD3 |
| :--- | ---: | ---: | ---: |
| y | 0.4027356 | 0.02967881 | 0.18839237 |
| smk | 0.1727459 | -0.06794271 | -0.09386294 |
| w | -0.5792456 | -0.16300221 | 0.07341123 |

- LD1 positive if y large and/or w small.
$>$ LD2 positive if w small.


## Discriminant scores for data

```
peanuts.2 %>% select(y, w, starts_with("x"))
```

|  | y | w | x.LD1 | x.LD2 | x.LD3 |
| :--- | ---: | ---: | ---: | ---: | ---: |
| 1 | 195.3 | 51.4 | 1.417354 | 1.01233393 | 0.26467918 |
| 2 | 194.3 | 53.7 | 2.204444 | -0.38421359 | -1.12526629 |
| 3 | 189.7 | 55.5 | -5.562217 | 1.10184441 | 0.78720394 |
| 4 | 180.4 | 44.4 | -6.056558 | 3.88530191 | -0.05263163 |
| 5 | 203.0 | 49.8 | 6.084370 | 1.25027629 | 1.25054957 |
| 6 | 195.9 | 45.8 | 7.131192 | 1.06649258 | -1.24422021 |
| 7 | 202.7 | 60.4 | 1.430084 | -1.11831802 | 1.09926555 |
| 8 | 197.6 | 54.1 | 2.282572 | 0.04938762 | 0.07958437 |
| 9 | 193.5 | 57.8 | -1.045438 | -0.85884902 | -0.67463274 |
| 10 | 187.0 | 58.6 | -4.022969 | -1.22292871 | -1.89677191 |
| 11 | 201.5 | 65.0 | -1.596806 | -1.95130266 | 1.14518230 |
| 12 | 200.0 | 67.2 | -2.266028 | -2.83002474 | 0.36705787 |

$\rightarrow$ Obs. 5 and 6 have most positive LD1: large y, small w.

- Obs. 4 has most positive LD2: small w.


## Plot LD1 vs. LD2, labelling by combo



## "Bi-plot" from ggbiplot

ggbiplot(peanuts.1, groups = factor(peanuts.combo\$combo))


## Installing ggbiplot

- ggbiplot not on CRAN, so usual install. packages will not work.
- Install package devtools first (once):
install.packages("devtools")
- Then install ggbiplot (once):
library (devtools)
install_github("vqv/ggbiplot")


## Cross-validation

- So far, have predicted group membership from same data used to form the groups - dishonest!
- Better: cross-validation: form groups from all observations except one, then predict group membership for that left-out observation.
- No longer cheating!
- Illustrate with peanuts data again.


## Misclassifications

- Fitting and prediction all in one go:

```
p <- lda(combo ~ y + smk + w,
    data = peanuts.combo, CV = TRUE)
peanuts. }3\mathrm{ <- cbind(peanuts.combo, class = p$class,
    posterior = p$posterior)
with(peanuts.3, table(obs = combo, pred = class))
    pred
\begin{tabular}{crrrrrr} 
obs & \(5 \_1\) & \(5 \_2\) & \(6 \_1\) & \(6 \_2\) & \(8 \_1\) & \(8 \_2\) \\
\(5 \_1\) & 0 & 0 & 0 & 2 & 0 & 0 \\
\(5 \_2\) & 0 & 1 & 0 & 0 & 1 & 0 \\
\(6 \_1\) & 0 & 0 & 2 & 0 & 0 & 0 \\
\(6 \_2\) & 1 & 0 & 0 & 1 & 0 & 0 \\
\(8 \_1\) & 0 & 1 & 0 & 0 & 0 & 1 \\
\(8 \_2\) & 0 & 0 & 0 & 0 & 0 & 2
\end{tabular}
```

- Some more misclassification this time.


## Repeat of LD plot



## Posterior probabilities

peanuts. 3 \% \% \%

```
mutate(across(starts_with("posterior"), \(p) round(p, 3))) %>%
select(combo, class, starts_with("posterior"))
```

combo class posterior.5_1 posterior.5_2 posterior.6_1 posterior.6_2

| 1 | $5 \_1$ | $6 \_2$ | 0.162 | 0.00 | 0.000 | 0.838 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| 2 | $5 \_1$ | $6 \_2$ | 0.200 | 0.00 | 0.000 | 0.799 |
| 3 | $5 \_2$ | $8 \_1$ | 0.000 | 0.18 | 0.000 | 0.000 |
| 4 | $5 \_2$ | $5 \_2$ | 0.000 | 1.00 | 0.000 | 0.000 |
| 5 | $6 \_1$ | $6 \_1$ | 0.194 | 0.00 | 0.669 | 0.137 |
| 6 | $6 \_1$ | $6 \_1$ | 0.000 | 0.00 | 1.000 | 0.000 |
| 7 | $6 \_2$ | $6 \_2$ | 0.325 | 0.00 | 0.000 | 0.667 |
| 8 | $6 \_2$ | $5 \_1$ | 0.821 | 0.00 | 0.000 | 0.179 |
| 9 | $8 \_1$ | $8 \_2$ | 0.000 | 0.00 | 0.000 | 0.000 |
| 10 | $8 \_1$ | $5 \_2$ | 0.000 | 1.00 | 0.000 | 0.000 |
| 11 | $8 \_2$ | $8 \_2$ | 0.001 | 0.00 | 0.000 | 0.004 |
| 12 | $8 \_2$ | $8 \_2$ | 0.000 | 0.00 | 0.000 | 0.000 |
|  | posterior.8_1 posterior.8_2 |  |  |  |  |  |
| 1 |  | 0.000 | 0.000 |  |  |  |
| 2 |  | 0.000 | 0.000 |  |  |  |
| 3 |  | 0.820 | 0.000 |  |  |  |
| 4 |  | 0.000 | 0.000 |  |  |  |
| 5 |  | 0.000 | 0.000 |  |  |  |

## Why more misclassification?

- When predicting group membership for one observation, only uses the other one in that group.
- So if two in a pair are far apart, or if two groups overlap, great potential for misclassification.
- Groups 5_1 and 6_2 overlap.
-5_2 closest to 8_1s looks more like an 8_1 than a 5_2 (other one far away).
-8_1s relatively far apart and close to other things, so one appears to be a 5_2 and the other an 8_2.


## Example 3: professions and leisure activities

>15 individuals from three different professions (politicians, administrators and belly dancers) each participate in four different leisure activities: reading, dancing, TV watching and skiing. After each activity they rate it on a $0-10$ scale.

- How can we best use the scores on the activities to predict a person's profession?
- Or, what combination(s) of scores best separate data into profession groups?


## The data

```
my_url <- "http://ritsokiguess.site/datafiles/profile.txt"
active <- read_delim(my_url, " ")
active
```



## Discriminant analysis

```
active.1 <- lda(job ~ reading + dance + tv + ski, data = active)
active.1
Call:
lda(job ~ reading + dance + tv + ski, data = active)
Prior probabilities of groups:
        admin bellydancer politician
    0.3333333 0.3333333 0.3333333
Group means:
            reading dance tv ski
admin 5.0 2.0 1.8 3.8
bellydancer 6.6 9.4 5.8 7.4
politician 5.0 4.8 5.2 5.0
Coefficients of linear discriminants:
            LD1 LD2
reading -0.01297465 -0.4748081
dance -0.95212396 -0.4614976
tv -0.47417264 1.2446327
ski 0.04153684 -0.2033122
Proportion of trace:
        LD1 LD2
0.8917 0.1083
```


## Comments

- Two discriminants, first fair bit more important than second.
- LD1 depends (negatively) most on dance, a bit on tv.
- LD2 depends mostly (negatively) on tv.


## Misclassification

```
p <- predict(active.1)
active.2 <- cbind(active, p)
with(active.2, table(obs = job, pred = class))
```

pred
obs
admin bellydancer politician
admin
bellydancer politician

0
0
0
0
5
Everyone correctly classified.

## Plotting LDs

$\mathrm{g}<-\operatorname{ggplot}($ active.2, aes $(\mathrm{x}=\mathrm{x} . \mathrm{LD} 1, \mathrm{y}=\mathrm{x} . \mathrm{LD} 2$, colour $=\mathrm{job}$, la geom_point() + geom_text_repel() + guides(colour = "none")
g


## Biplot

ggbiplot(active.1, groups = active\$job)


## Comments on plot

- Groups well separated: bellydancers top left, administrators top right, politicians lower middle.
- Bellydancers most negative on LD1: like dancing most.
- Administrators most positive on LD1: like dancing least.
- Politicians most negative on LD2: like TV-watching most.


## Plotting individual persons

Make label be identifier of person. Now need legend:

```
active.2 %>% mutate(person = row_number()) %>%
    ggplot(aes(x = x.LD1, y = x.LD2, colour = job,
        label = person)) +
    geom_point() + geom_text_repel()
```



## Posterinr nrohahilities

active. 2 \%>\% mutate(across(starts_with("posterior"), <br>(p) round(p, 3))) \%>\% select(job, class, starts_with("posterior"))

|  | job |  | class | posterior. admin |
| :--- | ---: | ---: | ---: | ---: | posterior.bellydancer

## posterior.politician

$\begin{array}{ll}1 & 0.000 \\ 2 & 0.000\end{array}$
$3 \quad 0.000$
$4 \quad 0.000$
$5 \quad 0.003$
6
0.997

## Cross-validating the jobs-activities data

Recall: no need for predict:

```
p <- lda(job ~ reading + dance + tv + ski, data = active,
active. }3\mathrm{ <- cbind(active, class = p$class, posterior = p$pc
with(active.3, table(obs = job, pred = class))
```

pred
obs admin bellydancer politician $\begin{array}{rrr}\text { admin } & \text { bellydancer } & \text { politician } \\ 5 & 0 & 0 \\ 0 & 4 & 1 \\ 0 & 0 & 5\end{array}$

This time one of the bellydancers was classified as a politician.

## and lonk at the nocterior nrohahilities

active. $3 \%>\%$
mutate(across(starts_with("posterior"), <br>(p) round(p, 3))) \%>\% select(job, class, starts_with("post"))
job class posterior.admin posterior.bellydancer
1 bellydancer bellydancer

| 0.000 | 1.000 |
| :--- | :--- |
| 0.000 | 1.000 |

$\begin{array}{ll}\text { bellydancer bellydancer } 0.000 & 1.000\end{array}$
$\begin{array}{lll}\text { bellydancer bellydancer } 0.000 & 1.000\end{array}$
$\begin{array}{lll}\text { bellydancer bellydancer } 0.000 & 1.000\end{array}$
$\begin{array}{lll}\text { bellydancer politician } 0.000 & 0.001\end{array}$
$\begin{array}{lll}\text { politician politician } & 0.006 & 0.000\end{array}$
$\begin{array}{lll}\text { politician politician } 0.001 & 0.000\end{array}$
$\begin{array}{lll}\text { politician politician } 0.000 & 0.000\end{array}$
$\begin{array}{lll}\text { politician politician } 0.000 & 0.009\end{array}$
$\begin{array}{lll}\text { politician politician } 0.000 & 0.000\end{array}$
$\begin{array}{lll}\text { admin admin } \quad 1.000 & 0.000\end{array}$
12 admin admin
13 admin admin
$\begin{array}{lll}14 & \text { admin } & \text { admin } \\ 15 & \text { admin } & \text { admin }\end{array}$
$\begin{array}{lll}14 & \text { admin } & \text { admin } \\ 15 & \text { admin } & \text { admin }\end{array}$
1.000
0.000
$1.000 \quad 0.000$
1.000
0.000
posterior.politician
0.819
0.000

$$
0.000
$$

$$
0.000
$$

$$
0.000
$$

$$
4
$$

$$
0.000
$$

5

## Comments

- Bellydancer was "definitely" a politician!
$>$ One of the administrators might have been a politician too.


## Why did things get misclassified?

Go back to plot of
discriminant scores:
$>$ one bellydancer much closer to the politicians,
$>$ one administrator a bit closer to the politicians.


## Example 4: remote-sensing data

- View 25 crops from air, measure 4 variables $\mathrm{x} 1-\mathrm{x} 4$.
- Go back and record what each crop was.
- Can we use the 4 variables to distinguish crops?


## The data

```
my_url <- "http://ritsokiguess.site/datafiles/remote-sensing.txt"
crops <- read_table(my_url)
crops %>% print(n = 25)
# A tibble: 25 x 6
\begin{tabular}{lrrrrl} 
crop & x1 & x2 & x3 & x4 cr \\
<chr> & <dbl> & <dbl> & <dbl> & <dbl> & <chr>
\end{tabular}
```


## Discriminant analysis

```
crops.1 <- lda(crop ~ x1 + x2 + x3 + x4, data = crops)
crops.1
```

Call:
lda(crop $\sim x 1+x 2+x 3+x 4$, data $=$ crops $)$
Prior probabilities of groups:
Corn Cotton Soybeans Sugarbeets
$\begin{array}{llll}0.28 & 0.24 & 0.24 & 0.24\end{array}$

Group means:

|  |  | x1 | x2 | x3 |
| :--- | ---: | ---: | ---: | ---: |$\quad$ x4

Coefficients of linear discriminants:
LD1 LD2 LD3
$\begin{array}{llll}\mathrm{x} 1 & 0.14077479 & 0.007780184 & -0.0312610362\end{array}$
$\begin{array}{llll}\mathrm{x} 2 & 0.03006972 & 0.007318386 & 0.0085401510\end{array}$
x3 -0.06363974-0.099520895-0.0005309869
x4-0.00677414-0.035612707 0.0577718649

Proportion of trace:
LD1 LD2 LD3
0.80440 .18320 .0124

## Assessing

$>3$ LDs (four variables, four groups).

- 1st two important.
- LD1 mostly x 1 (minus)
- LD2 x3 (minus)


## Predictions

$>$ Thus:

```
p <- predict(crops.1)
crops.2 <- cbind(crops, p)
with(crops.2, table(obs = crop, pred = class))
```

pred
obs Corn Cotton Soybeans Sugarbeets

| Corn | 6 | 0 | 1 | 0 |
| :--- | :--- | :--- | :--- | :--- |
| Cotton | 0 | 4 | 2 | 0 |
| Soybeans | 2 | 0 | 3 | 1 |
| Sugarbeets | 0 | 0 | 3 | 3 |

- Not very good, eg. only half the Soybeans and Sugarbeets classified correctly.


## Plotting the LDs

```
ggplot(crops.2, aes(x = x.LD1, y = x.LD2, colour = crop))
    geom_point()
```



Corn (red) mostly left, cotton (green) sort of right, soybeans and sugarbeets (blue and purple) mixed up.

## Biplot

## ggbiplot(crops.1, groups = crops\$crop)



## Comments

- Corn low on LD1 (left), hence low on x1
$>$ Cotton tends to be high on LD1 (high x1)
$>$ one cotton very low on LD2 (high x3?)
- Rather mixed up.


## Posterior probs (some)

```
crops.2 %>% mutate(across(starts_with("posterior"), \(p) round(p, 3))) %>%
    filter(crop != class) %>%
    select(crop, class, starts_with("posterior"))
```

|  | crop | class posterior.Corn posterior. Cotton posterior.Soybeans |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: |
| 4 | Corn | Soybeans | 0.443 | 0.034 | 0.494 |
| 11 | Soybeans | Sugarbeets | 0.010 | 0.107 | 0.299 |
| 12 | Soybeans | Corn | 0.684 | 0.009 | 0.296 |
| 13 | Soybeans | Corn | 0.467 | 0.199 | 0.287 |
| 15 | Cotton | Soybeans | 0.056 | 0.241 | 0.379 |
| 17 | Cotton | Soybeans | 0.066 | 0.138 | 0.489 |
| 20 | Sugarbeets | Soybeans | 0.381 | 0.395 |  |
| 21 | Sugarbeets | Soybeans | 0.106 | 0.518 |  |
| 24 | Sugarbeets | Soybeans | 0.088 | 0.489 |  |
|  | posterior.Sugarbeets |  |  |  |  |
| 4 |  | 0.029 |  |  |  |
| 11 |  | 0.584 |  |  |  |
| 12 |  | 0.011 |  |  |  |
| 13 |  | 0.047 |  |  |  |
| 15 |  | 0.324 |  |  |  |
| 17 |  | 0.306 |  |  |  |
| 20 |  | 0.078 |  |  |  |
| 21 |  | 0.232 |  |  |  |
| 24 |  | 0.216 |  |  |  |

## Comments

- These were the misclassified ones, but the posterior probability of being correct was not usually too low.
- The correctly-classified ones are not very clear-cut either.


## MANOVA

Began discriminant analysis as a followup to MANOVA. Do our variables significantly separate the crops?

```
response <- with(crops, cbind(x1, x2, x3, x4))
crops.manova <- manova(response ~ crop, data = crops)
summary(crops.manova)
```

|  | Df Pillai | approx F num Df den $\operatorname{Df}$ | $\operatorname{Pr}(>F)$ |  |
| :--- | ---: | :--- | ---: | :--- |
| crop | 3 | 0.9113 | 2.1815 | 12 | 60 | 0.02416 |
| :--- | :--- |

Residuals 21

Signif. codes: $0{ }^{\prime * * * '} 0.001$ '**' 0.01 '*' 0.05 '.' 0.1 '

## Box's M test

We should also run Box's $M$ test to check for equal variance of each variable across crops:
summary (BoxM(response, crops\$crop))
Box's M Test
Chi-Squared Value $=69.42634$, $\mathrm{df}=30$ and p -value: $5.79 \mathrm{e}-05$

- The P -value for the M test is smaller even than our guideline of 0.001 . So we should not take the MANOVA seriously.
- Apparently at least one of the crops differs (in means) from the others. So it is worth doing this analysis.
- We did this the wrong way around, though!


## The right way around

- First, do a MANOVA to see whether any of the groups differ significantly on any of the variables.
- Check that the MANOVA is believable by using Box's M test.
- If the MANOVA is significant, do a discriminant analysis in the hopes of understanding how the groups are different.
- For remote-sensing data (without Clover):
- LD1 a fair bit more important than LD2 (definitely ignore LD3).
- LD1 depends mostly on x 1, on which Cotton was high and Corn was low.

D Discriminant analysis in MANOVA plays the same kind of role that Tukey does in ANOVA.

